Multiplicity considerations for analyses of non-exchangeable subgroups

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MANY GOALS OF SUBGROUP ANALYSES

- \blacktriangleright Estimate average treatment effects within subgroups
- ▶ Estimate consistency or heterogeneity among subgroups
- \blacktriangleright Identify population with exceptional effects
- \blacktriangleright Identify population that benefits
- \blacktriangleright Ftc.

BENEFITING SUBGROUP IDENTIFICATION

- ▶ Personalized treatment effect (PTE) ∆(*x*) depending on covariate vector *x*, more positive is better
	- ▶ Could be CATE (E[*Y* ¹ *− Y* 0 *|x*]), hazard ratio, log odds ratio, etc.
- ▶ Benefiting subgroup B = *{x* : ∆(*x*) *> δ}*
- \blacktriangleright Want some estimate \widehat{B} with nice properties
- ▶ Types of patients only identified as far as *x* allows, i.e., no statements about individual units, e.g., $P[Y^1 > Y^0 \mid X \in B]$

NON-EXCHANGEABLE SUBGROUPS

- ▶ Each subpopulation indexed by a covariate point *x* may be considered a subgroup
	- ▶ Benefit: subgroups, thresholds, etc. need not be prespecified, just variables
	- ▶ Challenge: many subgroups! Potentially inconvenient shape of identified benefiting subgroup
- ▶ PTE correlated between similar *x*, *x ′* a priori (Bayesian) and as estimates
	- **Benefit:** accounting for positive correlation mitigates power loss from multiplicity adjustments
	- ▶ Challenge: difficult to get analytic results about joint distributions for multiplicity adjustments

REASONS FOR MULTIPLICITY CONTROL

- ▶ Need based on decisions being made
- \triangleright Regulatory: limit probability of marketing to populations not benefiting from treatment (FWER)

▶ P[$\widehat{B} \not\subseteq B$] < α , *not* P [$\int_{\widehat{B}} \Delta(x) dx \le \delta$] < α

- ▶ Cost-benefit thresholds: limit proportion of identified subgroup not benefiting from treatment (FDR)
	- ▶ Proportion of the $x \in \widehat{B}$ weighted by population distribution not satisfying ∆(*x*) *> δ*

FWER CONTROL

- ▶ FDA on multiple endpoints (draft guidance): FDA's concern for controlling the Type I error probability is to minimize the chances of a false favorable conclusion for any of the primary or secondary endpoints, regardless of which and how many endpoints in the study have no effect.
- ▶ Analogy to subgroups: Our concern for controlling the Type I error probability is to minimize the chances of a false favorable conclusion for any subgroups (here, *x*), regardless of which and how many subgroups in the study have no effect.

FWER CONTROL (CONTINUED)

- \triangleright Estimator \widehat{B} will collect **x** for which we are confident that $\Delta(x) > \delta$
- ▶ Controlling FWER means limiting probability that any $x \in \widehat{B}$ fails $\Delta(x) > \delta$
- Result: FWER is probability that \widehat{B} is not a subset of B.

CONFIDENCE/CREDIBLE SUBGROUPS

- ▶ Benefiting subgroup: types of patients who benefit from treatment
- ▶ Exclusive confidence subgroup: should contain *only* types of patients who benefit
- ▶ Inclusive confidence subgroup: should contain *all* types of patients who benefit

VISUALIZATION

FREQUENTIST DEFINITION

Definition (Confidence subgroups)

If $S_B(\theta)$ is the benefiting subgroup, then α -level inclusive confidence subgroups *SI*(*Y*) and exclusive confidence subgroups *SE*(*Y*) are subsets of the population such that under repeated sampling:

 $P_Y[S_E(Y) \subseteq S_B(\theta) \subseteq S_I(Y)] \geq 1 - \alpha$

INTERPRETATION

INVERTING SIMULTANEOUS INTERVAL ESTIMATES

SIMULTANEOUS INTERVAL ESTIMATES

- ▶ *M*-draw Monte Carlo estimate of sampling distribution of estimated effect surface $\widehat{\Delta}(x)$ approximately normal
- \blacktriangleright Restrict covariate space to C
- \triangleright Find the 1 α restricted-space confidence band:

$$
\widehat{\Delta}(x) \pm W^*_{\alpha,\mathrm{C}}\sqrt{\frac{1}{M-1}\sum_{m=1}^{M}\left\{\widehat{\Delta}^{(m)}(x)-\widehat{\Delta}(x)\right\}^2}
$$

where $W_{\alpha,\mathsf{C}}^*$ is the 1 α quantile of the distribution of

$$
W_C^{(m)} = \sup_{\mathbf{x} \in \mathcal{C}} \frac{\left| \widehat{\Delta}^{(m)}(\mathbf{x}) - \widehat{\Delta}(\mathbf{x}) \right|}{\sqrt{\frac{1}{M-1} \sum_{m=1}^{M} \left\{ \widehat{\Delta}^{(m)}(\mathbf{x}) - \widehat{\Delta}(\mathbf{x}) \right\}^2}}
$$

WEIGHTED FDR/PPV CONTROL

\blacktriangleright Measure μ on covariate space

- \blacktriangleright E.g., population covariate distribution
- \blacktriangleright Could be estimated by empirical distribution
- \blacktriangleright FDR control:

$$
\mathsf{E}\left[\frac{\mu(\widehat{\mathbf{B}}\setminus\mathbf{B})}{\mu(\widehat{\mathbf{B}})}\right]<\alpha
$$

► Positive predictive value analogue: 1 – Ε $\left[\frac{\mu(\widehat{B}\setminus B)}{\mu(\widehat{B})}\right]$ μ (B) i is PPV of "diagnostic test" assigning patients to benefiting subgroup

ESTIMATED FDR FOR INCLUSION IN \widehat{B}

- ▶ Yekutieli & Benjamini, "Resampling-based false discovery rate controlling multiple test procedures for correlated test statistics." *JSPI* 82(1-2): 171–196, 1999.
- **Bootstrap**

$$
Z^{(k)}(\mathbf{x}_i) = \left\{ \widehat{\Delta}^{(k)}(\mathbf{x}_i) - \widehat{\Delta}(\mathbf{x}_i) \right\} / \sqrt{\sum_{k=1}^K \left\{ \widehat{\Delta}^{(k)}(\mathbf{x}_i) - \widehat{\Delta}(\mathbf{x}_i) \right\}^2}
$$

▶ Unadjusted bootstrap *p*-values from *Z* scores

▶ For fixed *p* estimate

$$
Q_{\widehat{B}}(p) = \frac{\mu(\widehat{B}(p) \setminus B)}{\mu(\widehat{B}(p) \setminus B) + \mu(\widehat{B}(p) \cap B)},
$$
\n(1)

the weighted FDR of the testing procedure which rejects the null hypotheses with *p*-values less than *p*

SIMULATION

Figure: N = 300, varying number of uniform binary covariates, effect is 1 iff first two covariates are 1. Target FDR is 10%. Res-True: resampling with true weights; Res-Emp: resampling with empirical weights; BH: Benjamini-Hochberg FDR procedure; FWER: FWER-controlling confidence subgroups.

SIMULATION

Table: $N = 300$; thin plate spline fits (mgcv); target 10% for respective error rates

- ▶ Resampling procedures conservative when effect not uniformly null
- ▶ Asymptotically, points with ∆(*x*) *> δ* are correctly identified almost surely
- ▶ Res-True very conservative for quadr. case (investigating)

EXAMPLE

- ▶ Data from sequence of four clinical trials of Alzheimer's disease treatments with same placebo control and SoC active control
- ▶ Endpoint: change in ADAS-Cog disease severity over 12 weeks
- ▶ Covariates: sex, APoE4 carrier status, baseline disease severity, rate of decline from diagnosis to baseline

EXAMPLE FWER

Figure: FWER contours. The exclusive confidence subgroup controlling FWER at or below 5% is the region within the 0*.*05 contour, etc.

EXAMPLE FDR

Figure: FDR contours. The exclusive confidence subgroup controlling FDR at or below 5% is the region within the 0*.*05 contour, etc.

CONCLUSION

- \triangleright Benefiting subgroups can be built up from atomic subgroups (e.g., covariate points)
- \blacktriangleright Multiplicity control goals have useful intuitive interpretations in terms of benefiting subgroups (*FWER*: estimator contained in benefiting subgroup, 1 *− FDR*: PPV of assignment)
- ▶ Bayesian or resampling methods allow flexible model choice decoupled from multiplicity handling
- ▶ Building subgroups from non-exchangeable atoms benefits from correlation between similar covariate points

SELECTED REFERENCES

- 1. Schnell. "Monte Carlo approaches to frequentist multiplicity-adjusted benefiting subgroup identification." *Stat Methods Med Res* 2020.
- 2. Schnell, Tang, Müller, Carlin. "Semiparametric benefiting subgroup identification via credible subgroups." *Clin Trials* 2018.
- 3. FDR-controlling method to be submitted to *Clin Trials* special issue.
- 4. credsubs R package on CRAN (FWER only for now) email: schnell.31@osu.edu